
Principles of Communications

Lecture 10: Stochastic Processes

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Outlines

- Terminology of Random Processes
- Correlation and Power Spectral Density
- Linear Systems and Random Processes
- Narrowband Noise

Random (Stochastic) Processes

- ***Informal Definition:*** The outcomes (events) of a chance (prob) experiment are mapped into *functions of time* (waveforms).

Cf.: Random variables: outcomes are mapped to *numbers*.

Note: Each waveform is called a **sample function**, or a **realization**. The totality of all sample functions is called an **ensemble**.

- Random Process ~ RP
- We next show the axiomatic-theory approach.

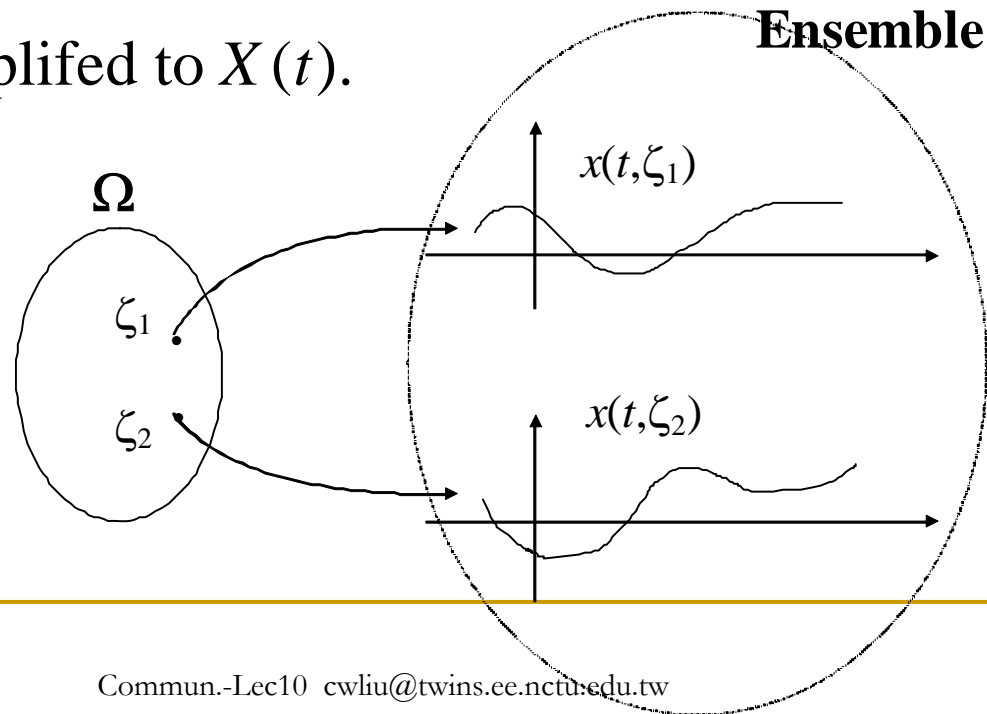
Formal Definition

$(\Omega, \mathfrak{F}, P)$: a probability space

Random Process: $X(t, \zeta)$ is a mapping from Ω to an ensemble of sample functions and for each fixed t , $X(t, \zeta)$ is a random variable.

Note: For each fixed ζ , $X(t, \zeta)$ is NOT random, it is an ordinary (deterministic) time function.

$X(t, \zeta)$ is often simplified to $X(t)$.



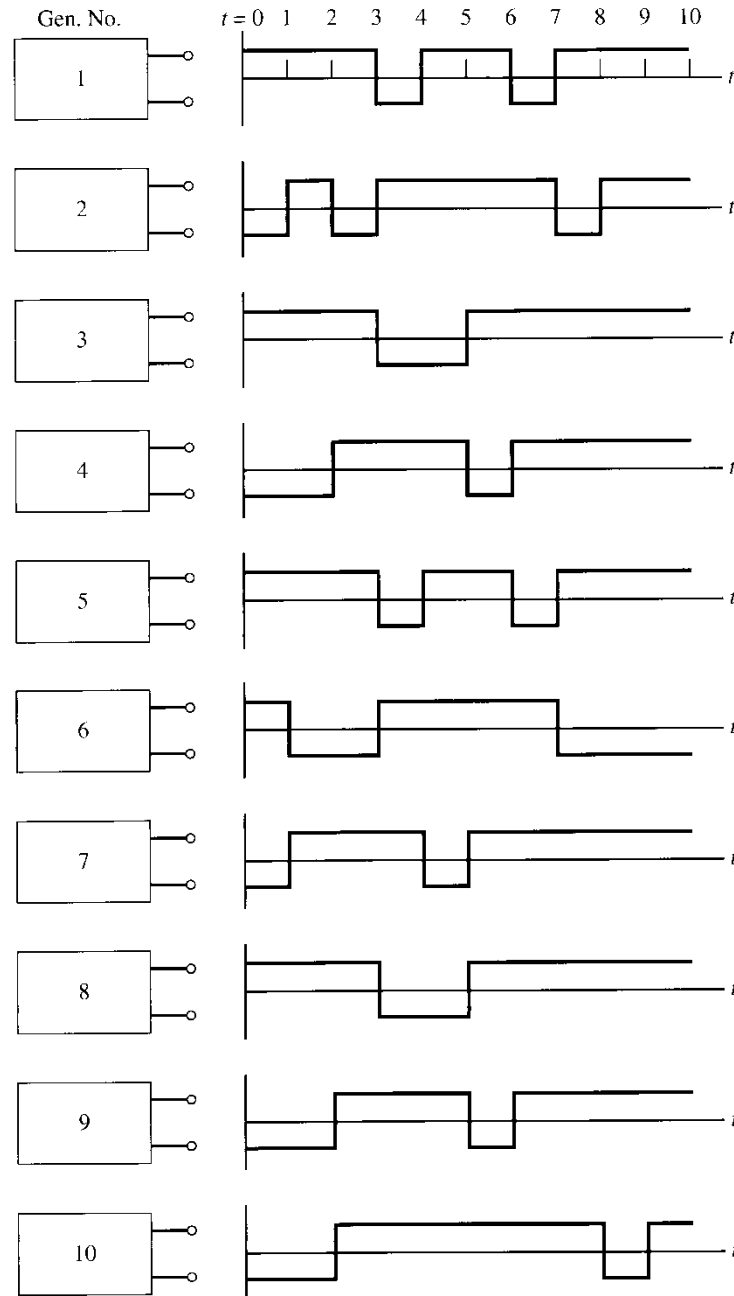


Figure 6.1
A statistically identical set of binary waveform generators with typical outputs.

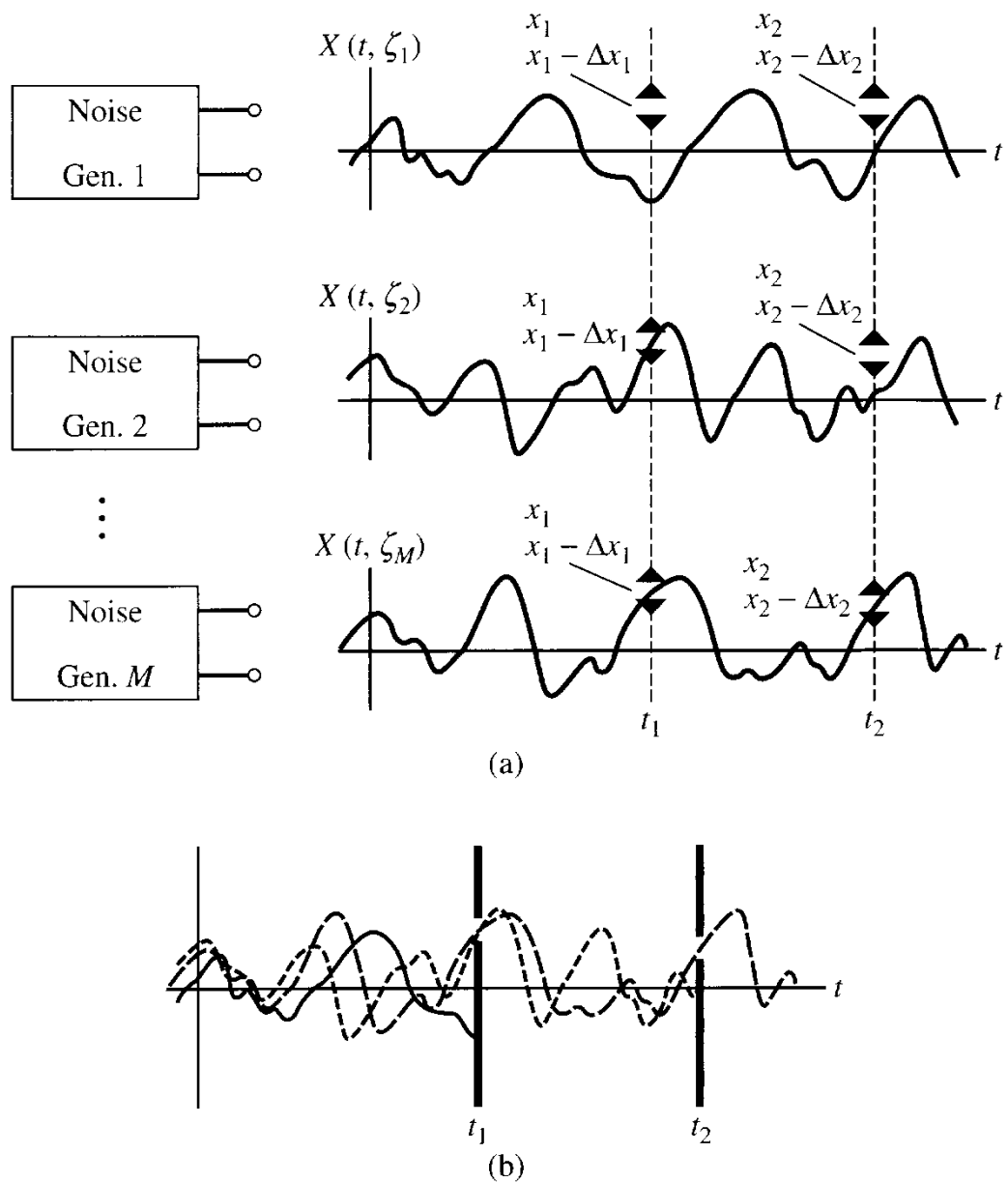


Figure 6.2

Typical sample functions of a random process and illustration of the relative-frequency interpretation of its joint pdf. (a) Ensemble of sample functions. (b) Superposition of the sample functions shown in (a).

Statistical Description

A random process is statistically specified if we know ALL its n th order joint pdf's for any t_1, t_2, \dots, t_n , for any n .

$$\begin{aligned} & f(x_1, t_1; x_2, t_2; \dots; x_n, t_n) dx_1 dx_2 \dots dx_n \\ &= P(x_1 - dx_1 < X_1 \leq x_1; x_2 - dx_2 < X_2 \leq x_2; \dots; \\ & \quad x_n - dx_n < X_n \leq x_n). \end{aligned}$$

Note: Strictly speaking, all the n th order joint pdf's are not sufficient to determine a random process completely.

For example, we need to impose a continuity requirement on the sample functions.

Ensemble Average

- Statistics based on prob model -- Expectation

The result is a deterministic value or function (non-random)

Mean function: $m_x(t) \equiv \bar{X}(t) \equiv E\{X(t)\} = \int_{-\infty}^{\infty} \alpha f_X(\alpha; t) d\alpha.$

Variance function: $\sigma_x^2(t) \equiv E\{[X(t) - \bar{X}(t)]^2\}.$

(Auto)correlation function:

$$\begin{aligned} R_x(t_1, t_2) &\equiv E\{X(t_1)X^*(t_2)\} \\ &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \alpha_1 \alpha_2 f_{x_1 x_2}(\alpha_1, t_1; \alpha_2, t_2) d\alpha_1 d\alpha_2. \end{aligned}$$

Time Average

Stochastic integral: integration over TIME

\approx sum of r.v.'s = r.v. Let $\Delta t_i = t_i - t_{i-1}$

$$I \equiv \int_{T_1}^{T_2} X(t) dt \equiv \lim_{n \rightarrow \infty} \sum_{i=1}^n X(t_i) \Delta t_i, \quad T_1 \leq t_0 \leq t_1 \leq \dots \leq t_n \leq T_2.$$

$$\text{Time average: } \langle X(t) \rangle \equiv \lim_{T \rightarrow \infty} \frac{1}{2T} \int_{-T}^T X(t) dt.$$

$$\langle [X(t) - \langle X(t) \rangle]^2 \rangle \equiv \lim_{T \rightarrow \infty} \frac{1}{2T} \int_{-T}^T (X(t) - \langle X(t) \rangle)^2 dt.$$

$$\langle X(t) X^*(t + \tau) \rangle \equiv \lim_{T \rightarrow \infty} \frac{1}{2T} \int_{-T}^T X(t) X^*(t + \tau) dt.$$

Stationary RP

- (*Strict sense*) **Stationary**: (pdf is shift-invariant)

A random process $X(t)$ is statistically stationary if $X(t)$ has the same n th-order pdf as $X(t + \Gamma)$, for $\forall \Gamma$ and $\forall n > 0$.

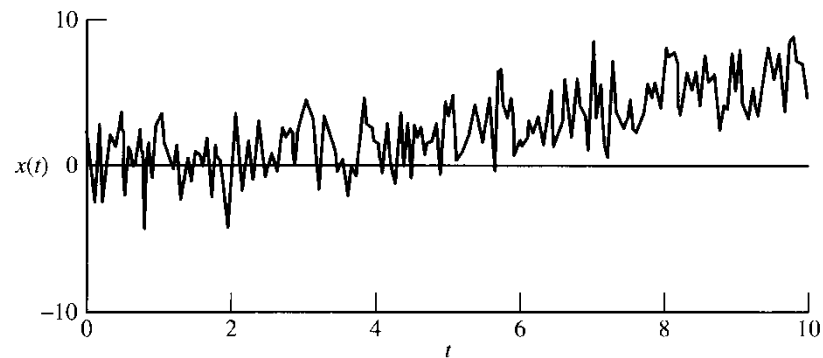
That is,

$$f(x_1, t_1; x_2, t_2; \cdots; x_n, t_n) = f(x_1, t_1 + T; x_2, t_2 + T; \cdots; x_n, t_n + T).$$

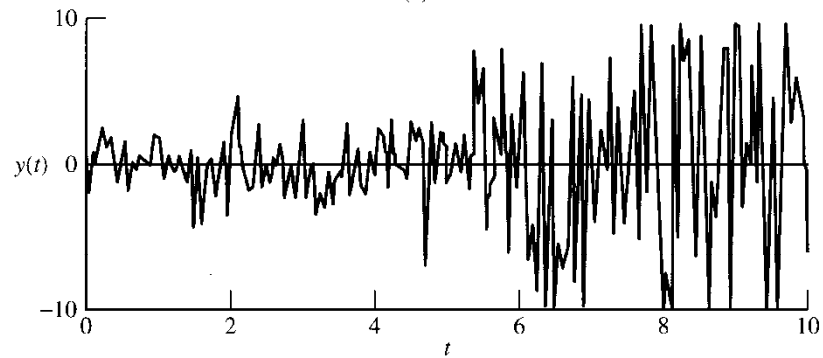
Remark: Consequently, the n th-order joint pdf depends only on the time differences $t_2 - t_1, t_3 - t_1, \cdots, t_n - t_1$.

Thus, $R_x(t_1, t_2) = R_x(0, t_2 - t_1) \equiv R_x(t_2 - t_1)$.

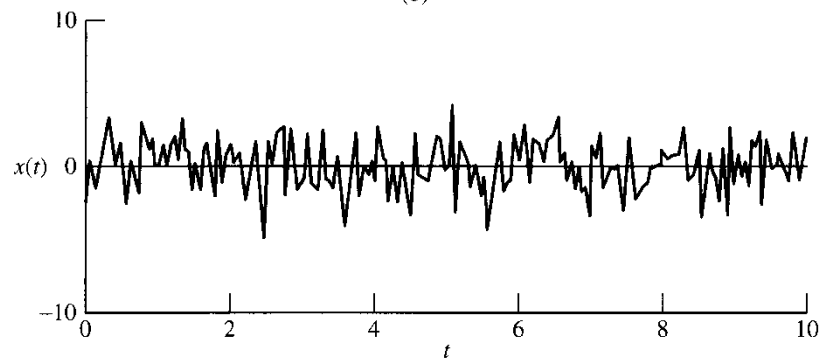
$$m_x(t) = \text{constant} = m_x(0).$$



(a)



(b)



(c)

Figure 6.3

Sample functions of nonstationary processes contrasted with a sample function of a stationary process.

(a) Time-varying mean. (b) Time-varying variance. (c) Stationary.

Wide-sense Stationary

- A random process $X(t)$ is **wide-sense stationary (WSS)** if $E[X(t)] = \text{constant}$, and
$$E[X(t)X^*(t+\tau)] = R_X(\tau) \quad \text{function of } \tau \text{ only.}$$

Remarks: (1) strict-sense stationary \Rightarrow wide-sense stationary

(2) For Gaussian random processes, WSS = strict-sense stationary

(A Gaussian process is completely specified by only two parameters: mean and covariance.)

Ergodicity

- **Ergodicity:** “Ensemble average = Time average”
- 3 types: in mean, in correlation, “strict-sense”
- Definition: **Ergodic in the mean:**

A WSS random process $X(t)$ is ergodic in the mean if

$$\langle X(t) \rangle = m_x = \text{constant.}$$

Remark: $\langle X(t) \rangle$ is the time average (a random variable).

A more rigorous definition:

$$M_T \equiv \frac{1}{2T} \int_{-T}^T X(t) dt \underset{T \rightarrow \infty}{=} m_x \text{ (in mean square sense)}$$

- A random variable Y_T converges to Y in m.s. sense, if $E[|Y_T - Y|^2] \rightarrow 0$, as $T \rightarrow \infty$.

Therefore, ergodic in the mean $\Leftrightarrow \lim_{T \rightarrow \infty} \text{var}\{M_T\} = 0$.

Ergodicity (2)

- Definition: **Ergodic in correlation**:

A WSS random process $X(t)$ is ergodic in correlation if

$$\lim_{T \rightarrow \infty} \left[\frac{1}{2T} \int_{-T}^T X(t + \tau) X^*(t) dt \right] \underset{m.s.}{=} R_X(\tau), \forall \tau.$$

- Definition: **Ergodic**

A random process $X(t)$ is ergodic if all time and ensemble averages are interchangeable. (The n th order moments of both are identical for $\forall n$.)

Remark: Ergodicity \Rightarrow strict-sense stationarity

Ex.: Ergodic RP

Ergodic process: $n(t) = A \cos(\omega_0 t + \Theta)$.

$$\Theta: \text{random variable} \quad f_{\Theta}(\theta) = \begin{cases} \frac{1}{2\pi}, & |\theta| \leq \pi \\ 0, & \text{otherwise} \end{cases}.$$

<Ensemble average> Mean: $\overline{n(t)} = E[A \cos(\omega_0 t + \Theta)]$

$$= \int_{-\infty}^{\infty} A \cos(\omega_0 t + \theta) f_{\Theta}(\theta) d\theta = \int_{-\pi}^{\pi} A \cos(\omega_0 t + \theta) \frac{1}{2\pi} d\theta = 0.$$

$$\text{Variance: } \overline{n^2(t)} = \int_{-\pi}^{\pi} (A \cos(\omega_0 t + \theta))^2 \frac{1}{2\pi} d\theta = \frac{A^2}{2}.$$

$$n(t) = A \cos(\omega_0 t + \Theta) \dots$$

<Time average>

$$\langle n(t) \rangle = \lim_{T \rightarrow \infty} \frac{1}{2T} \int_{-T}^T A \cos(\omega_0 t + \theta) dt = 0, \quad \text{for any } \theta.$$

$$\langle n^2(t) \rangle = \lim_{T \rightarrow \infty} \frac{1}{2T} \int_{-T}^T A^2 \cos^2(\omega_0 t + \theta) dt = \frac{A^2}{2}, \quad \text{for any } \theta.$$

$\therefore \langle n(t) \rangle = E[n(t)], \langle n^2(t) \rangle = E[n^2(t)] \Rightarrow n(t)$ is, at least, ergodic in power. It is in fact ergodic.

Moreover, it is ergodic in autocorrelation function.

Ensemble: $R_X(\tau) = E[X(t)X^*(t + \tau)]$

$$= E[A \cos(\omega_0 t + \Theta) A \cos(\omega_0 t + \omega_0 \tau + \Theta)]$$

$$= \frac{A^2}{2} \{ E[\cos(2\omega_0 t + 2\Theta + \omega_0 \tau)] + E[\cos(\omega_0 \tau)] \}$$

$$= \frac{A^2}{2} \left\{ \cos(\omega_0 \tau) + \int_{-\pi}^{\pi} \cos(2\omega_0 t + \omega_0 \tau + 2\theta) \frac{1}{2\pi} d\theta \right\}$$

$$= \frac{A^2}{2} \cos(\omega_0 \tau).$$

Time: $\langle X(t)X^*(t + \tau) \rangle = \lim_{T \rightarrow \infty} \frac{1}{2T} \int_{-T}^T A^2 \cos(\omega_0 t + \theta) \cos(\omega_0 t + \omega_0 \tau + \theta) dt$

$$= \lim_{T \rightarrow \infty} \frac{A^2}{2T} \int_{-T}^T \left[\frac{1}{2} \cos(\omega_0 \tau) + \frac{1}{2} \cos(2\omega_0 t + \omega_0 \tau + 2\theta) \right] dt$$

$$= \frac{A^2}{2} \cos(\omega_0 \tau) = R_X(\tau).$$

Ex.: Non-stationary RP

- Nonstationary process (nonergodic) $n(t) = A \cos(\omega_0 t + \Theta)$.

Θ : a random variable with $f_{\Theta}(\theta) = \begin{cases} \frac{2}{\pi}, & |\theta| \leq \frac{\pi}{4} \\ 0, & \text{otherwise} \end{cases}$.

$$\text{Mean: } \overline{n(t)} = \int_{-\frac{\pi}{4}}^{\frac{\pi}{4}} A \cos(\omega_0 t + \theta) \frac{2}{\pi} d\theta = \frac{2A}{\pi} \sin(\omega_0 t + \theta) \Big|_{-\frac{\pi}{4}}^{\frac{\pi}{4}} = \frac{2\sqrt{2}A}{\pi} \cos \omega_0 t.$$

$$\begin{aligned} \text{Variance: } \overline{n^2(t)} &= \int_{-\frac{\pi}{4}}^{\frac{\pi}{4}} (A \cos(\omega_0 t + \theta))^2 \frac{2}{\pi} d\theta \\ &= \int_{-\frac{\pi}{4}}^{\frac{\pi}{4}} \frac{A^2}{\pi} [1 + \cos(2\omega_0 t + 2\theta)] d\theta = \frac{A^2}{2} + \frac{A^2}{\pi} \cos 2\omega_0 t. \end{aligned}$$

Random Telegraph Waveform

Definition: (i) At $t = t_0$, $\Pr\{X(t_0) = A\} = 0.5 = \Pr\{X(t_0) = -A\}$

(ii) The number of switching (K) in any time interval T

□ Poisson distribution, i.e.,

$$P_T(K) = \frac{(\alpha T)^K}{K!} e^{-\alpha T}, \text{ for some } \alpha > 0.$$

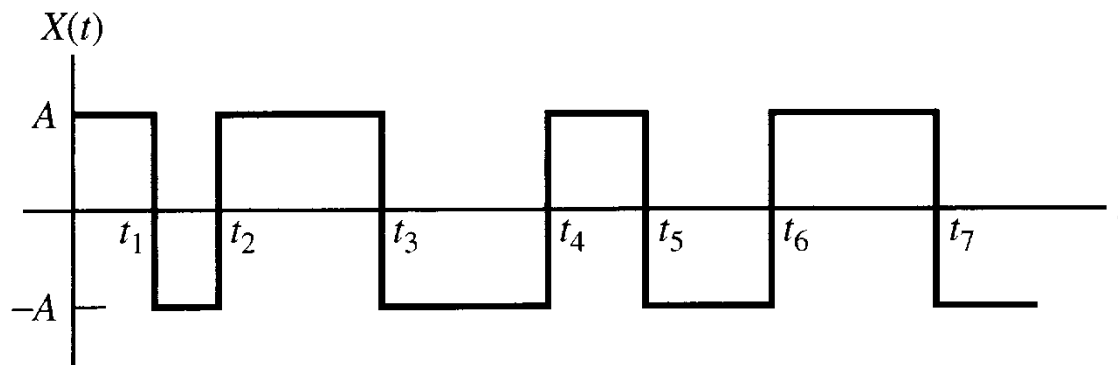


Figure 6.4

Sample function of a random telegraph waveform.

Mean of $X(t) = 0$. ($\because X(t) = \pm A$, equally likely)

Correlation function: $R_X(\tau) = E[X(t)X(t + \tau)]$

$$= A^2 \cdot P\{X(t) \text{ and } X(t + \tau) \text{ have the same sign}\}$$

$$- A^2 \cdot P\{X(t) \text{ and } X(t + \tau) \text{ have different signs}\}$$

$$= A^2 \cdot P\{\text{even number of switchings in } (t, t + \tau)\}$$

$$- A^2 \cdot P\{\text{odd number of switchings in } (t, t + \tau)\}$$

$$= A^2 \sum_{\substack{k=0, \\ k \text{ even}}}^{\infty} \frac{(\alpha\tau)^k}{k!} e^{-\alpha\tau} - A^2 \sum_{\substack{k=0 \\ k \text{ odd}}}^{\infty} \frac{(\alpha\tau)^k}{k!} e^{-\alpha\tau}$$

$$\sum_{k=0}^{\infty} \frac{(-\alpha\tau)^k}{k!} = e^{-\alpha\tau}$$

$$= A^2 e^{-\alpha\tau} \sum_{k=0}^{\infty} \frac{(-\alpha\tau)^k}{k!} = A^2 e^{-2\alpha\tau}, \text{ for } \alpha\tau > 0 \text{ (or } \tau > 0 \text{).}$$

Similarly, for $\tau < 0$, $R_X(\tau) = A^2 e^{2\alpha\tau}$. $\therefore R_X(\tau) = A^2 e^{-2\alpha|\tau|}$.

Interpretation of Moments

- For ergodic PR, the moments (statistics) have the *physical* meanings.

1. Mean $\overline{X(t)} = \langle X(t) \rangle$ is the dc component.

2. $\overline{X(t)^2} = \langle X(t) \rangle^2$ is the dc power.

3. 2nd moment: $\overline{X^2(t)} = \langle X^2(t) \rangle$ is the total power.

4. Variance: $\sigma_X^2 = \overline{X(t)^2} - \overline{X(t)}^2 = \langle X^2(t) \rangle - \langle X(t) \rangle^2$

is the power in the ac (time-varying) component.

5. The total power $\overline{X^2(t)} = \sigma_X^2 + \overline{X(t)}^2$ is the ac plus dc power.

Remark: The time averages can be measured in the laboratory.

Power Spectral Density (PSD)

For a WSS process $X(t)$, the PSD is defined as

$$S_X(f) \equiv FT[R_X(\tau)] \text{ (Fourier transform)}$$

- What is the relationship between the PSD and the Fourier transform of a realization of the random process?

$$\text{Let } n_T(t, \zeta_i) = \begin{cases} n(t, \zeta_i), & |t| < \frac{T}{2} \\ 0, & \text{otherwise} \end{cases}.$$

$$N_T(f, \zeta_i) = FT\{n_T(t, \zeta_i)\} = \int_{-\frac{T}{2}}^{\frac{T}{2}} n(t, \zeta_i) e^{-j\omega t} dt.$$

Note: In general, $n(t, \zeta_i)$ has infinite energy. Hence, $FT[n(t, \zeta_i)]$ does not exist.

With finite T , for each ζ_i , energy spectral density of $n(t, \zeta_i) = |N_T(f, \zeta_i)|^2$.

$$\text{Power spectral density of } n(t, \zeta_i) = \frac{|N_T(f, \zeta_i)|^2}{T}.$$

- For all ζ_i , let $S_n(f) = \lim_{T \rightarrow \infty} \frac{1}{2T} \overline{|N_T(f, \zeta_i)|^2}$.

- **The Wiener - Khinchin theorem:**

$$S_n(f) = \lim_{T \rightarrow \infty} \frac{1}{2T} E\{|FT[n_{2T}(t)]|^2\}.$$

- **Proof:** $|FT[n_{2T}(t)]|^2 = \int_{-T}^T n(t) e^{-j\omega t} dt \int_{-T}^T n^*(\sigma) e^{j\omega\sigma} d\sigma$
 $= \int_{-T}^T \int_{-T}^T n(t) n^*(\sigma) e^{-j\omega(t-\sigma)} dt d\sigma.$

$$E\{|FT[n_{2T}(t)]|^2\} = \int_{-T}^T \int_{-T}^T E\{n(t) n^*(\sigma)\} e^{-j\omega(t-\sigma)} dt d\sigma$$

$$= \int_{-T}^T \int_{-T}^T R_n(t-\sigma) e^{-j\omega(t-\sigma)} dt d\sigma.$$

Change of variables: $u = t - \sigma$ and $v = t$.

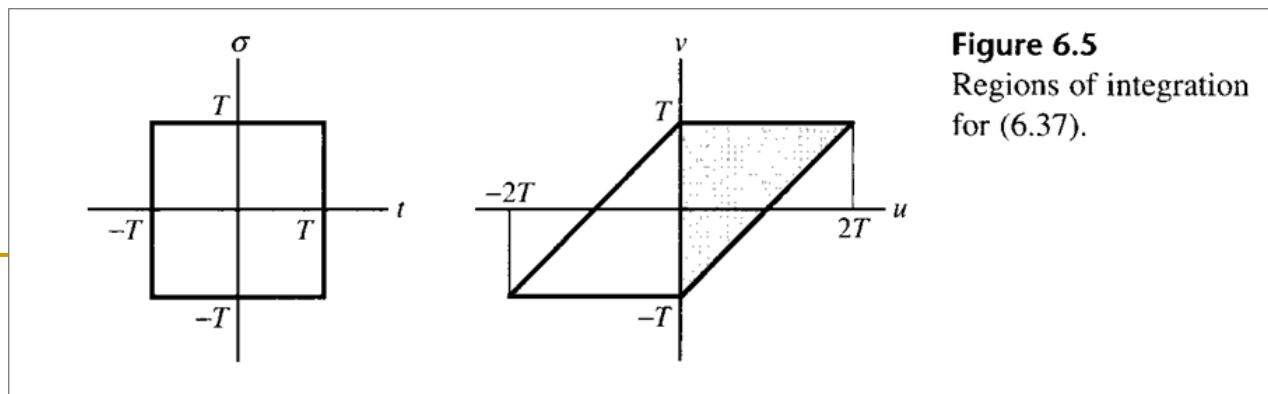


Figure 6.5
Regions of integration
for (6.37).

For a given t value, the range of $u : \{t - T, t + T\}$.

For a given u value,

if $u > 0$, the range of $v : \{u - T, T\}$,

else ($u < 0$), the range of $v : \{-T, u + T\}$.

$$\begin{aligned} E\{|FT[n_{2T}(t)]|^2\} &= \int_{-2T}^0 \left(\int_{-T}^{u+T} R_n(u) e^{-j\omega u} dv \right) du \\ &\quad + \int_0^{2T} \left(\int_{u-T}^T R_n(u) e^{-j\omega u} dv \right) du \\ &= \int_{-2T}^0 R_n(u) e^{-j\omega u} \left(\int_{-T}^{u+T} dv \right) du + \int_0^{2T} R_n(u) e^{-j\omega u} \left(\int_{u-T}^T dv \right) du \\ &= \int_{-2T}^0 (2T + u) R_n(u) e^{-j\omega u} du + \int_0^{2T} (2T - u) R_n(u) e^{-j\omega u} du \\ &= 2T \int_{-2T}^{2T} \left(1 - \frac{|u|}{2T}\right) R_n(u) e^{-j\omega u} du. \end{aligned}$$

$$\begin{aligned} \lim_{T \rightarrow \infty} \frac{1}{2T} E\{|FT[n_{2T}(t)]|^2\} &= \lim_{T \rightarrow \infty} \int_{-2T}^{2T} \left(1 - \frac{|u|}{2T}\right) R_n(u) e^{-j\omega u} du \\ &= \int_{-\infty}^{\infty} R_n(u) e^{-j\omega u} du = FT\{R_n(u)\}. \end{aligned}$$

Ex.: Random Phase Sinusoidal Wave

Sinusoidal wave with random phase:

$$n(t) = A \cos(\omega_0 t + \Theta). \quad \Theta : \text{uniform pdf } [-\pi, \pi]$$

$$R_n(\tau) = \frac{A^2}{2} \cos \omega_0 \tau.$$

$$S_n(f) = \frac{A^2}{4} [\delta(f + f_0) + \delta(f - f_0)].$$

Properties of $R(\tau)$

- $R(\tau)$: Autocorrelation function of a WSS random process

(1) $R(0) \geq |R(\tau)|, \forall \tau.$

Proof: $E\{[X(t) \pm X(t + \tau)]^2\} \geq 0$

$$\Rightarrow \overline{X^2(t)} \pm 2\overline{X(t)X(t + \tau)} + \overline{X^2(t + \tau)} \geq 0$$

$$\Rightarrow 2R(0) \pm 2R(\tau) \geq 0$$

$$\Rightarrow -R(0) \leq R(\tau) \leq R(0).$$

- (2) $R(\tau)$ is even; $R(\tau) = R(-\tau)$ (if $x(t)$ is real).

Poof: From the definition (of WSS).

(3) $\lim_{|\tau| \rightarrow \infty} R(\tau) = \overline{X(t)}^2$, if $X(t)$ does not contain a periodic component.

Proof: If $X(t)$ is not periodic, $X(t)$ and $X(t + \tau)$ are nearly independent when $\tau \rightarrow \infty$.

Thus, $\lim_{|\tau| \rightarrow \infty} \overline{X(t)X(t + \tau)} \approx \overline{X(t)}\overline{X(t + \tau)} = \overline{X(t)}^2$.

(4) If $X(t)$ is periodic with period T_0 , then $R(\tau)$ is periodic with the same period. [$R(\tau) = R(T_0 + \tau)$]

(5) $S(f) = FT[R(\tau)] \geq 0, \forall f$.

Proof: By Wiener-Khinchin theorem,

$$S(f) = \lim_{T \rightarrow \infty} \frac{1}{2T} E\{|FT[X_{2T}(t)]|^2\} \geq 0.$$

Properties of $S(f)$

(1) $S(f) = FT[R(\tau)] \geq 0, \forall f.$

(2) $S(f)$ is real-valued. [$\because R(\tau)$ is conjugate symmetric]

(3) If $X(t)$ is real, $S(f)$ is even. [$\because R(\tau)$ is real]

(4) "dc power" = $\int_{-\infty}^{\infty} R(\tau) d\tau = S(0).$

"total power" = $\int_{-\infty}^{\infty} S(f) df = R(0).$

Note : $S(f)$ is the average power at frequency f .

Crosscorrelation

- For two RPs: $X(t)$ and $Y(t)$

- Definition: **Cross - correlation**:

$$R_{XY}(t_1, t_2) \equiv E\{X(t_1)Y^*(t_2)\}.$$

If $X(t)$ and $Y(t)$ are jointly WSS,

$$R_{XY}(\tau) \equiv E\{X(t)Y^*(t + \tau)\}.$$

- Definition: **Cross - power spectral density**:

If $X(t)$ and $Y(t)$ are joint WSS, $S_{XY}(f) \equiv \mathfrak{F}\{R_{XY}(\tau)\}.$

• Definition: For two random processes $X(t)$ and $Y(t)$,

(1) **Uncorrelated**: if $R_{XY}(t_1, t_2) = m_X(t_1)m_Y^*(t_2), \forall t_1, t_2$.

(2) **Orthogonal**: if $R_{XY}(t_1, t_2) = 0, \forall t_1, t_2$.

(3) **Independent**: if $f_{XY}(x_1, y_1, t_1; x_2, y_2, t_2; \dots; x_n, y_n, t_n)$
 $= f_X(x_1, t_1; x_2, t_2; \dots; x_n, t_n) f_Y(y_1, t_1; y_2, t_2; \dots; y_n, t_n)$.

<Remarks>

(1) Independent \Rightarrow Uncorrelated.

(2) Uncorrelated $\Rightarrow (X(t) - m_{X(t)})$ and $(Y(t) - m_{Y(t)})$ are orthogonal.

(3) {Uncorrelated and either $m_{X(t)} = 0$ or $m_{Y(t)} = 0$ } \Rightarrow orthogonal.

(4) Uncorrelated and Gaussian \Rightarrow Independent.